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## Statistically Correcting for Chance using the Adjusted and Standardized Mutual Information Measures

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## Definition of Mutual Information

Mutual Information (MI) quantifies the information shared between two categorical random variables $X$ and $Y$ :

$$
\begin{aligned}
\operatorname{MI}(X, Y) & =\sum_{x \in \mathcal{X}} \sum_{y \in \mathcal{Y}} p_{X, Y}(x, y) \log \frac{p_{X, Y}(x, y)}{p_{X}(x) p_{Y}(y)} \\
& =H(Y)-H(Y \mid X)
\end{aligned}
$$

where $H$ is the entropy function which quantifies uncertainty. MI intuitively quantifies the uncertainty of $Y$ explained by $X^{1}$.
Characteristics

- $\operatorname{MI}(X, Y)=0$ if $X$ and $Y$ are independent;
- MI is maximized when one variable is a deterministic function of the other. E.g. $Y=f(X) \Rightarrow \operatorname{MI}(X, Y)=H(Y)$.

[^0]
## Extension to continuous random variables

MI can also quantify the dependency between two continuous random variables:

$$
\operatorname{MI}(X, Y)=\int_{-\infty}^{+\infty} \int_{-\infty}^{+\infty} f_{X, Y}(x, y) \log \frac{f_{X, Y}(x, y)}{f_{X}(x) f_{Y}(y)}
$$

## Characteristics

- $\operatorname{MI}(X, Y)=0$ if $X$ and $Y$ are independent;


## Importance of MI

MI is a compelling tool to assess the strength of the dependency between features because it is based on a well-established theory and quantifies non-linear interactions which might be missed if e.g. the Pearson's correlation coefficient $r(X, Y)$ is used.

## Definition

## Estimation of MI

## Categorical variables

The estimation for the categorical case is straightforward: the empirical probability distribution for $p_{X, Y}(x, y), p_{X}(x)$, and $p_{Y}(y)$ is computed on data and plugged in the MI formula. In this case, MI is also a linear function of the $G$-statistics used in likelihood-ratio tests : $G=2 N \cdot \mathrm{MI}$ with $N$ number of records.

## Continuous variables

A number of different estimators have been proposed for MI in the continuous case. The standard approach consists in discretizing the space of possible values for $X$ and $Y$. There are also many possible approaches for discretization [Garcia et al., 2013], however the straightforward way is to discretize $X$ and $Y$ according to equal-width or equal-frequency binning.

| Group | Type | Citation |
| :--- | :--- | :---: |
| Discretization | Discretization equal width | [Steuer et al., 2002] |
|  | Discretization equal frequency | [teuer et al., 2002] |
|  | Adaptive Discretization | [Cellucci et al., 2005] |
| Others | Nearest Neighbour | [Kraskov et al., 2004] |
|  | Kernel Density Estimation | [Moon et al., 1995] |

Table: List of possible estimators.

## Non-exhaustive list of other dependency measures

Information theory gave birth to some new dependency measures (also based on discretization) in the last few years:

| Acronym | Name | Citation |
| :--- | :--- | :---: |
| MIC | Maximal Information Coefficient | [Reshef et al., 2011] |
| GMIC | Generalized Mean Information Coefficient | [Luedthe and Tran, 2013] |
| MID | Mutual Information Dimension | [Sugiyama and Borgwardt, 2013] |

Of course the number of possible non-linear dependency measures in use is large:

| Acronym | Name | Citation |
| :--- | :--- | :---: |
| dCorr | Distance Correlation | [Székely et al., 2009] |
| RDC | Randomized Dependency Coefficient | [Lopez-Paz et al., 2013] |
| HSIC | Hilbert-Schmidt Independence Criterion | [Gretton et al., 2005] |

However, information theory provides a well-established framework and it has been successfully employed for a variety of applications...

## Applications

## Supervised data mining

- Feature selection [Nguyen et al., 2014b, Nguyen et al., 2014a];
- Decision tree induction [Criminisi et al., 2012].


## Unsupervised data mining

- External clustering validation [Romano et al., 2014];
- Generation of alternative or multi-view clusterings [Dang and Bailey, 2015, Müller et al., 2013];
- The exploration of the clustering space using results from the Meta-Clustering algorithm [Caruana et al., 2006].


## Exploratory data mining

- Analysis of neural time-series data [Cohen, 2014];
- Reverse engineering of biological networks [Villaverde et al., 2013];


## Application examples

## Remark:

In the rest of the talk we focus on MI for categorical variables or the discretized version of continuous variables.

## Examples:

To gain intuition about MI computation we describe in detail 2 application examples:

1. External clustering validation;
2. Decision tree induction.

## Application example (1): external clustering validation

Task: Compare a clustering solution B to a reference clustering A.

Example
$N=15$ data points
reference clustering A with 2 clusters, stars $\hat{\sim}$ and circles $O$


## Application example (1): external clustering validation

Task: Compare a clustering solution B to a reference clustering A.

Example
$N=15$ data points
reference clustering A with 2 clusters, stars $\hat{\sim}$ and circles $O$


## Applications

## MI computed on a contingency table

MI is estimated on data via a contingency table that assess the amount of overlap between $\mathbf{A}$ and $\mathbf{B}$


## MI computation

MI between the two clusterings $\mathbf{A}$ and B is computed on a contingency table $\mathcal{M}$ using the empirical probability distributions $\frac{n_{i j}}{N}, \frac{a_{i}}{N}$, and $\frac{b_{j}}{N}$ :

$$
\operatorname{MI}(\mathbf{A}, \mathbf{B})=\sum_{i=1}^{r} \sum_{j=1}^{c} \frac{n_{i j}}{N} \log \frac{n_{i j} N}{a_{i} b_{j}}
$$



Contingency table $\mathcal{M}$
$a_{i}=\sum_{j} n_{i j}$ are the row marginals and $b_{j}=\sum_{i} n_{i j}$ are the column marginals.

## Applications

## Application example (2): decision tree induction

Task: Find the most informative feature $\mathbf{F}$ to the target class $\mathbf{C}$.
$\mathrm{MI}(\mathbf{F}, \mathbf{C})$ is still computed on a contingency table. In this scenario MI is also known as the Information Gain: $\mathrm{IG}(\mathbf{F}, \mathbf{C})=\mathrm{MI}(\mathbf{F}, \mathbf{C})$
E.g. if the class $\mathbf{C}=$ cancer and a feature $\mathbf{F}=$ smoker.


## Applications

## Limitations

MI is a well-established tool to compare two random variables but it is has some limitations that can be overcome by its statistical adjustments.

## Limitation and solution

- Non-intuitive range of variation
$\Rightarrow$ Solution: the Normalized Mutual Information (NMI) [Kvalseth, 1987]; Ensure the range of the measure is in the range $[0,1]$
- Non-zero baseline
$\Rightarrow$ Solution: the Adjusted Mutual Information (AMI) [Vinh et al., 2009]; Value of measure is expected to be zero when sampling at random features to be correlated.
- Selection bias
$\Rightarrow$ Solution: the Standardized Mutual Information (SMI) [Romano et al., 2014]; Avoid preferring features with many bins/categories.


## Mutual Information

# Normalized Mutual Information 

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## Motivation

## Definition of the Normalized Mutual Information

## Limitation of MI

MI has a non-intuitive range of variation. What does an MI of 5.6 mean ?

## Solution

MI can be normalized by its maximum value in order to vary in the interval $[0,1]$ :

$$
\mathrm{NMI}=\frac{\mathrm{MI}}{\max \mathrm{Ml}}
$$

Many possible upper bounds for $\operatorname{MI}(\mathbf{A}, \mathbf{B})$ :
$\min \{H(\mathbf{A}), H(\mathbf{B})\} \leq \sqrt{H(\mathbf{A}) \cdot H(\mathbf{B})} \leq \frac{1}{2}(H(\mathbf{A})+H(\mathbf{B})) \leq \max \{H(\mathbf{A}), H(\mathbf{B})\} \leq H(\mathbf{A}, \mathbf{B})$
Depending on the chosen upper bound, it is possible to obtain information theoretic distance measures with metric properties [Vinh et al., 2010]. A distance measure with metric properties is indeed useful for designing efficient algorithms that exploit the nice geometric properties of metric spaces [Meilă, 2012].

## Motivation

## Normalization of Mutual Information

In [Vinh et al., 2010] we propose a review of possible normalization choices for MI.

Table: Normalization of Mutual Information.

| Name | Expression | Range | Related sources |
| :---: | :---: | :---: | :---: |
| $\mathrm{NMI}_{\text {joint }}$ | $\frac{\mathrm{Ml}(\mathbf{A}, \mathbf{B})}{\mathrm{H}(\mathbf{A}, \mathbf{B})}$ | $[0,1]$ | $[$ Yao, 2003] |
| $\mathrm{NMI}_{\text {max }}$ | $\frac{\mathrm{Ml}(\mathbf{A}, \mathbf{B})}{\max \{\mathrm{H}(\mathbf{A}), \mathrm{H}(\mathbf{B})\}}$ | $[0,1]$ | $[$ Kvalseth, 1987] |
| $\mathrm{NMI}_{\text {sum }}$ | $\frac{2 \mathrm{Ml}(\mathbf{A}, \mathbf{B})}{\mathrm{H}(\mathbf{A})+\mathrm{H}(\mathbf{B})}$ | $[0,1]$ | [Kvalseth, 1987] |
| $\mathrm{NMI}_{\text {sqrt }}$ | $\frac{\mathrm{Ml}(\mathbf{A}, \mathbf{B})}{\sqrt{H(\mathbf{A}) H(\mathbf{B})}}$ | $[0,1]$ | [Strehl and Ghosh, 2002] |
| $\mathrm{NMI}_{\text {min }}$ | $\frac{\mathrm{Ml}(\mathbf{A}, \mathbf{B})}{\min \{H(\mathbf{A}), H(\mathbf{B})\}}$ | $[0,1]$ |  |

Table: Distance measures based on MI.

| Name | Expression | Range | Metric | Related sources |
| :---: | :--- | :---: | :---: | :---: |
| $D_{\text {joint }}(V I)$ | $H(\mathbf{A}, \mathbf{B})-\mathrm{MI}(\mathbf{A}, \mathbf{B})$ | $[0, \log N]$ | $\checkmark$ | [Yao, 2003] |
| (Variation of Information $)$ | $\max \{H(\mathbf{A}), H(\mathbf{B})\}-\mathrm{MI}(\mathbf{A}, \mathbf{B})$ | $[0, \log N]$ | $\checkmark$ |  |
| $D_{\max }$ | $\frac{1}{2}[H(\mathbf{A})+H(\mathbf{B})]-\mathrm{MI}(\mathbf{A}, \mathbf{B})$ | $[0, \log N]$ | $\checkmark$ |  |
| $D_{\text {sum }}\left(\equiv \frac{1}{2} D_{\text {joint }}\right)$ | $\sqrt{H(\mathbf{A}) H(\mathbf{B})}-\mathrm{MI}(\mathbf{A}, \mathbf{B})$ | $[0, \log N]$ | $\boldsymbol{x}$ |  |
| $D_{\text {sqrt }}$ | $\min \{H(\mathbf{A}), H(\mathbf{B})\}-\mathrm{MI}(\mathbf{A}, \mathbf{B})$ | $[0, \log N]$ | $\boldsymbol{x}$ |  |
| $D_{\min }$ |  |  |  |  |

## Successful applications and limitations

NMI has been shown to be successful in:

- Clustering comparisons scenarios [Strehl and Ghosh, 2003, Wu et al., 2009];
- Decision tree induction [Quinlan, 1993];
- Feature selection [Estévez et al., 2009].

However NMI has some limitations
NMI does not have constant $\mathbf{0}$ baseline value for independent variables $\mathbf{A}$ and $\mathbf{B}$.

## Limitations

## Limitation on case study: external clustering validation

Task: Compare a clustering solution B to reference clustering A.
Experiment
$N=500$ data points
A with 10 clusters


Figure: If the clustering solution $\mathbf{B}$ is generated independently from $\mathbf{A}$ at random with $c$ clusters the average value of MI and NMI increases at the increase of the number of clusters.

Needs of statistical correction for MI

## Limitations

## Little affect of other approaches:

A correction for MI has already been proposed a while ago [Miller, 1955]:

$$
\mathrm{MI}(\text { Miller correction })=\mathrm{MI}-\frac{(r-1)(c-1)}{2 N}
$$

with $r, c$ number of bins and $N$ number of records.
However it seems not effective in the general case:


Figure: Clustering solutions B generated independently from A. Miller correction is not effective.

To address this issue we propose to statistically adjust MI for chance

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## The Adjusted Mutual Information

## Limitation of NMI

MI and NMI have non-zero baseline.

## Solution

Statistically adjust MI by the subtraction of its expected value under the null hypothesis of independence. The Adjusted Mutual Information (AMI) is defined as [Vinh et al., 2009]:

$$
\mathrm{AMI}=\frac{\mathrm{MI}-E[\mathrm{MI}]}{\max \mathrm{MI}-E[\mathrm{MI}]}
$$

The resulting measure is statistically normalized: it is equal to 0 when MI is equal to the expected value obtained by chance.

## Adjustment for chance

We compute the expected value of MI under the null hypothesis of independent clusterings $\mathbf{A}$ and $\mathbf{B}$.
we make use of the permutation model to compute it analytically: the distribution of MI is computed using all possible contingency tables $\mathcal{M}$ obtained by permutations.

## Motivation

## Expected Value

$E[\mathrm{MI}]$ is obtained by summation over all possible contingency tables $\mathcal{M}$ obtained by permutations.

$$
E[\mathrm{MI}]=\sum_{\mathcal{M}} \operatorname{MI}(\mathcal{M}) P(\mathcal{M})=\sum_{\mathcal{M}} \sum_{i, j} \frac{n_{i j}}{N} \log \frac{n_{i j} N}{a_{i} b_{j}} P(\mathcal{M})
$$

- No method to exhaustively generate $\mathcal{M}$
- extremely time expensive ( permutations $\mathcal{O}(n!)$ )

However, it is possible to swap the inner summation with the outer summation:

- $n_{i j}$ has a known hypergeometric distribution,
- Computation time dramatically reduced!


## Motivation

According to the different upper bound to MI used we obtain different versions of the Adjusted Mutual Information (AMI):

Table: Adjusted Mutual Information [Vinh et al., 2010].

| Name | Expression | Range |
| :---: | :---: | :---: |
| $\mathrm{AMI}_{\text {max }}$ | $\frac{\mathrm{Ml}(\mathbf{A}, \mathbf{B})-E[\mathrm{Ml}(\mathbf{A}, \mathbf{B})]}{\max \{H(\mathbf{A}), H(\mathbf{B})\}-E[\mathrm{Ml}(\mathbf{A}, \mathbf{B})]}$ | $[0,1]^{*}$ |
| $\mathrm{AMI}_{\text {sum }}$ | $\frac{\mathrm{Ml}(\mathbf{A}, \mathbf{B})-E[\mathrm{MI}(\mathbf{A}, \mathbf{B})]}{\frac{1}{2}(H(\mathbf{A})+H(\mathbf{B})-E[\mathrm{Ml}(\mathbf{A}, \mathbf{B})]}$ | $[0,1]^{*}$ |
| $\mathrm{AMI}_{\text {sqrt }}$ | $\frac{\mathrm{MI}(\mathbf{A}, \mathbf{B})-E[\mathrm{MI}(\mathbf{A}, \mathbf{B})]}{\sqrt{H(\mathbf{A}) \cdot H(\mathbf{B})}-E[\mathrm{MI}(\mathbf{A}, \mathbf{B})]}$ | $[0,1]^{*}$ |
| $\mathrm{AMI}_{\text {min }}$ | $\frac{\mathrm{Ml}(\mathbf{A}, \mathbf{B})-E[\mathrm{MI}(\mathbf{A}, \mathbf{B})]}{\min \{H(\mathbf{A}), H(\mathbf{B})\}-E[\mathrm{MI}(\mathbf{A}, \mathbf{B})]}$ | $[0,1]^{*}$ |
| *These measures are normalized in a statistical sense. |  |  |

## Speed considerations

The computational complexity of NMI depends just on the number of clusters:

$$
\mathcal{O}(r c)
$$

The computational complexity of AMI is linear in the number of records $N$ :

$$
\mathcal{O}(\max \{r N, c N\})
$$

However

- Useful when the number of data points is small because

$$
\lim _{N \rightarrow+\infty} E[\mathrm{MI}]=0
$$

- Somebody has recently parallelized it [Schmidt et al., 2014].


## Motivation

## Successful application

Task: Compare a clustering solution B to reference clustering A.
Experiment
$N=500$ data points
A with 10 clusters


Figure: AMI obtains 0 baseline when clusterings $\mathbf{B}$ are generated at random.

## Successful applications and limitations

AMI is becoming a popular tool to compare clusterings.

| Titie 1-20 | Cited by | Year |
| :---: | :---: | :---: |
| Information theoretic measures for clusterings comparison: is a correction for chance necessary? <br> NX Vinh, J Epps, J Balley <br> Proceedings of the 26th Annual International Conference on Machine Learning ... | 198 | 2009 |
| $\cdots$ |  |  |
| Information theoretic measures for clusterings comparison: Variants, properties, normalization and correction for chance <br> NX Vinh, J Epps, J Bailey <br> The Journal of Machine Learning Research 11, 2837-2854 | 159 | 2010 |

Figure: AMI is a polar tool for clustering comparisons.

However even AMI has some limitations:
AMI is affected by selection bias.

## Limitation on case study: selection of clustering solution

Task: Select the most similar clustering solution B to a reference clustering A.
Experiment
$N=500$ data points
A with 10 clusters
Each $\mathbf{B}$ is generated independently from $\mathbf{A}$ :

## Limitation on case study: selection of clustering solution

Task: Select the most similar clustering solution B to a reference clustering A.
Experiment
$N=500$ data points
A with 10 clusters
Each $\mathbf{B}$ is generated independently from $\mathbf{A}$ :

- One clustering solution $\mathbf{B}$ on $c=2$ clusters


## Limitation on case study: selection of clustering solution

Task: Select the most similar clustering solution B to a reference clustering A.
Experiment
$N=500$ data points
A with 10 clusters
Each $\mathbf{B}$ is generated independently from $\mathbf{A}$ :

- One clustering solution $\mathbf{B}$ on $c=2$ clusters
- One clustering solution $B$ on $c=6$ clusters


## Limitation on case study: selection of clustering solution

Task: Select the most similar clustering solution B to a reference clustering A.
Experiment
$N=500$ data points
A with 10 clusters
Each $\mathbf{B}$ is generated independently from $\mathbf{A}$ :

- One clustering solution $\mathbf{B}$ on $c=2$ clusters
- One clustering solution $\mathbf{B}$ on $c=6$ clusters
- One clustering solution B on $c=10$ clusters


## Limitation on case study: selection of clustering solution

Task: Select the most similar clustering solution B to a reference clustering A.
Experiment
$N=500$ data points
A with 10 clusters
Each $\mathbf{B}$ is generated independently from $\mathbf{A}$ :

- One clustering solution $\mathbf{B}$ on $c=2$ clusters
- One clustering solution $\mathbf{B}$ on $c=6$ clusters
- One clustering solution B on $c=10$ clusters
- One clustering solution B on $c=14$ clusters


## Limitation on case study: selection of clustering solution

Task: Select the most similar clustering solution B to a reference clustering A.
Experiment
$N=500$ data points
A with 10 clusters
Each $\mathbf{B}$ is generated independently from $\mathbf{A}$ :

- One clustering solution $\mathbf{B}$ on $c=2$ clusters
- One clustering solution $\mathbf{B}$ on $c=6$ clusters
- One clustering solution B on $c=10$ clusters
- One clustering solution B on $c=14$ clusters
- One clustering solution B on $c=18$ clusters


## Limitation on case study: selection of clustering solution

Task: Select the most similar clustering solution B to a reference clustering A.
Experiment
$N=500$ data points
A with 10 clusters
Each $\mathbf{B}$ is generated independently from $\mathbf{A}$ :

- One clustering solution $\mathbf{B}$ on $c=2$ clusters
- One clustering solution $\mathbf{B}$ on $c=6$ clusters
- One clustering solution B on $c=10$ clusters
- One clustering solution B on $c=14$ clusters
- One clustering solution B on $c=18$ clusters
- One clustering solution B on $c=22$ clusters


## Limitation on case study: selection of clustering solution

Task: Select the most similar clustering solution B to a reference clustering A.
Experiment
$N=500$ data points
A with 10 clusters
Each $\mathbf{B}$ is generated independently from $\mathbf{A}$ :

- One clustering solution $\mathbf{B}$ on $c=2$ clusters
- One clustering solution $\mathbf{B}$ on $c=6$ clusters
- One clustering solution B on $c=10$ clusters
- One clustering solution B on $c=14$ clusters

Select the $\mathbf{B}$ that yields the maximum $\operatorname{MI}(\mathbf{A}, \mathbf{B})$

Give a win to the solution that gets the highest value

- One clustering solution B on $c=18$ clusters
- One clustering solution B on $c=22$ clusters


## Limitation on case study: selection of clustering solution

Task: Select the most similar clustering solution B to a reference clustering A.
Experiment
$N=500$ data points
A with 10 clusters
Each $\mathbf{B}$ is generated independently from $\mathbf{A}$ :

- One clustering solution $\mathbf{B}$ on $c=2$ clusters
- One clustering solution $\mathbf{B}$ on $c=6$ clusters
- One clustering solution B on $c=10$ clusters
- One clustering solution B on $c=14$ clusters
- One clustering solution B on $c=18$ clusters

Select the $\mathbf{B}$ that yields the maximum $\operatorname{MI}(\mathbf{A}, \mathbf{B})$

Give a win to the solution that gets the highest value

REPEAT

- One clustering solution B on $c=22$ clusters


## Limitations

## Selection Bias

MI unfairly selects more often the solution with $c=22$ clusters.


## Limitations

## Also AMI is affected by selection bias

$$
\mathrm{AMI}=\frac{\mathrm{MI}-E[\mathrm{MI}]}{\sqrt{H(\mathbf{A}) \cdot H(\mathbf{B})}-E[\mathrm{MI}]}
$$



Also AMI is affected by selection bias

$$
\mathrm{AMI}=\frac{\mathrm{MI}-E[\mathrm{MI}]}{\sqrt{H(\mathbf{A}) \cdot H(\mathbf{B})}-E[\mathrm{MI}]}
$$



We have to take into account full distributional properties of MI: we proceed by subtracting its expected value and dividing by its standard deviation:
we propose to statistically standardize MI

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## Motivation

## Non-standardized variance

## Limitation of AMI

$\mathrm{MI}, \mathrm{NMI}$, and AMI are affected by selection bias.

## Solution

This behaviour is due to the non-standardized variance of AMI $\Rightarrow$ need of standardization.


Figure: AMI values have bigger variation when the number of clusters $c$ for $\mathbf{B}$ is high.

## Motivation

## Definition of Standardized Mutual Information

The Standardized Mutual Information (SMI) is defined as [Romano et al., 2014]:

$$
\mathrm{SMI}=\frac{\mathrm{MI}-E[\mathrm{MI}]}{\sqrt{\operatorname{Var}(\mathrm{MI})}}
$$

where we compute the expected value and the variance of Mutual Information under the null hypothesis of independent clusterings $\mathbf{A}$ and $\mathbf{B}$.

The SMI value is the number of standard deviations the mutual information is away from the expected value.

As in [Vinh et al., 2009] we make use of the permutation model to compute the expected value and the variance:
$\Rightarrow$ The distribution of MI is computed using all possible contingency tables $\mathcal{M}$ obtained by permutations.

## Variance Computation

We have to compute MI's second moment:

$$
\begin{aligned}
E\left[\mathrm{MI}^{2}\right] & =\sum_{\mathcal{M}} \operatorname{MI}(\mathcal{M})^{2} P(\mathcal{M})=\sum_{\mathcal{M}}\left(\sum_{i=1}^{r} \sum_{j=1}^{c} \frac{n_{i j}}{N} \log \frac{n_{i j} N}{a_{i} b_{j}}\right)^{2} P(\mathcal{M}) \\
& =\underbrace{\sum_{\mathcal{M}} \sum_{i, j, i^{\prime}, j^{\prime}} \frac{n_{i j}}{N} \log \frac{n_{i j} N}{a_{i} b_{j}} \cdot \frac{n_{i^{\prime} j^{\prime}}}{N} \log \frac{n_{i^{\prime} j^{\prime}} N}{a_{i^{\prime}} b_{j^{\prime}}} P(\mathcal{M})}_{\text {to swap }} \\
& =\underbrace{\sum_{i, j, i^{\prime}, j^{\prime}} \sum_{n_{i j}} \sum_{n_{i^{\prime} j^{\prime}}} \frac{n_{i j}}{N} \log \frac{n_{i j} N}{a_{i} b_{j}} \cdot \frac{n_{i^{\prime} j^{\prime}}}{N} \log \frac{n_{i^{\prime} j^{\prime}} N}{a_{i^{\prime}} b_{j^{\prime}}} P\left(n_{i j}, n_{i^{\prime} j^{\prime}}\right)}_{\text {swapped }}
\end{aligned}
$$

## Variance Computation

We have to compute MI's second moment:

$$
\begin{aligned}
E\left[\mathrm{MI}^{2}\right] & =\sum_{\mathcal{M}} \mathrm{MI}(\mathcal{M})^{2} P(\mathcal{M})=\sum_{\mathcal{M}}\left(\sum_{i=1}^{r} \sum_{j=1}^{c} \frac{n_{i j}}{N} \log \frac{n_{i j} N}{a_{i} b_{j}}\right)^{2} P(\mathcal{M}) \\
& =\underbrace{\sum_{\mathcal{M}} \sum_{i, j, i^{\prime}, j^{\prime}} \frac{n_{i j}}{N} \log \frac{n_{i j} N}{a_{i} b_{j}} \cdot \frac{n_{i^{\prime} j^{\prime}}}{N} \log \frac{n_{i^{\prime} j^{\prime}} N}{a_{i^{\prime}} b_{j^{\prime}}} P(\mathcal{M})}_{\text {to swap }} \\
& =\underbrace{\sum_{i, j, i^{\prime}, j^{\prime}} \sum_{n_{i j}} \sum_{n_{i^{\prime} j^{\prime}}} \frac{n_{i j}}{N} \log \frac{n_{i j} N}{a_{i} b_{j}} \cdot \frac{n_{i^{\prime} j^{\prime}}}{N} \log \frac{n_{i^{\prime} j^{\prime}} N}{a_{i^{\prime}} b_{j^{\prime}}} P\left(n_{i j}, n_{i^{\prime} j^{\prime}}\right)}_{\text {swapped }}
\end{aligned}
$$

Contribution: $P\left(n_{i j}, n_{i^{\prime} j^{\prime}}\right)$ computation is technically challenging.
We use the hypergeometric model: drawings from a urn with $N$ marbles with 3 colors, red, blue, and white.

## Bias Towards More Clusters Correction

MI and AMI unfairly select more often the solution with $c=22$ clusters:



## Characteristics of standardized measures

## Bias Towards Fewer Data Points Correction

## Reference clustering A on $N=100$ data points with 4 clusters

B induced independently on $N=20,40,60,80,100$ data points with 4 clusters.



人

## Unification property

The ability to compute a variance term allows extension of the existing measures:

- Variation of Information
- $G$-statistic


## Definitions:

$$
\mathrm{SVI}=\frac{E[\mathrm{VI}]-\mathrm{VI}}{\sqrt{\operatorname{Var}(\mathrm{VI})}}, \quad \mathrm{S} G=\frac{G-E[G]}{\sqrt{\operatorname{Var}(G)}}
$$

Theorem: The standardization unifies information theoretic measures:

$$
\mathrm{SMI}=\mathrm{SVI}=\mathrm{SG}
$$

## Speed considerations

The computational complexity of SMI is dominated by the computational complexity of $E\left[\mathrm{MI}^{2}\right]$ :

$$
\mathcal{O}\left(\max \left\{r c N^{3}, c^{2} N^{3}\right\}\right)
$$

However

- Useful when the number of data points is small;
- Faster than using the full distribution (compared to the $p$-value for the Fisher's exact test);
- Easily parallelizable.

Time in seconds for $4 \times 4$ tables with $N$ records

|  | 100 | 150 | 200 | 250 | 300 | 350 |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: |
| SMI | 0.65 | 1.53 | 2.94 | 5.00 | 7.59 | 11.00 |
| SMI (4 cores) | 0.30 | 0.51 | 0.97 | 1.52 | 2.33 | 3.35 |
| Fisher's | 0.65 | 11.32 | 242.67 | 844.62 | $\mathrm{~N} / \mathrm{A}$ | $\mathrm{N} / \mathrm{A}$ |

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## Summary

We discussed some enhancements to mutual information obtained by statistical correction for chance.

Limitation and solution

- Non-intuitive range of variation
$\Rightarrow$ Solution: the Normalized Mutual Information (NMI) [Kvalseth, 1987];
- Non-zero baseline
$\Rightarrow$ Solution: the Adjusted Mutual Information (AMI) [Vinh et al., 2009];
- Selection bias
$\Rightarrow$ Solution: the Standardized Mutual Information (SMI) [Romano et al., 2014];


## Take Away Message

Each variant is useful in some specific scenarios and there is a trade-off in computational complexity:


| Name | Range | Computational complexity |
| :--- | :--- | :--- |
| NMI | $[0,1]^{*}$ | $\mathcal{O}(r c)$ |
| AMI | $[0,1]$ | $\mathcal{O}(\max \{r N, c N\})$ |
| SMI | $[0, \infty)$ | $\mathcal{O}\left(\max \left\{r c N^{3}, c^{2} N^{3}\right\}\right)$ |
|  | $*$ non statistically normalized |  |

Table: Complexity when comparing two clusterings $\mathbf{A}$ and $\mathbf{B}$ with $r$ and c clusters on $N$ records.

## Open issues

There is a number of open issues for SMI:

- SMI achieves strength toward selection bias at the loss of normalization in the range $[0,1]$
$\Rightarrow$ need of statistical adjustment which allows normalization;
- SMI computational complexity might be problematic
$\Rightarrow$ at the large number of records $N, G$-statistic ( $G=2 N \cdot \mathrm{MI}$ ) can be approximated with a $\chi^{2}$ distribution. Need to find the scenarios where an exact SMI can be substituted by an approximation;
- SMI counts the number of standard deviations of MI, it might act as an exact $p$-value for MI. p-values quantifies the statistical significance of MI and this might sometimes interfere with the effect size of MI.
E.g. $\mathrm{SMI}=25.4$ ( 25.4 standard deviations away from mean). Is this closer to an effect size or an assessment of statistical significance ?
$\Rightarrow$ need of trade-offs between importance of statistical significance and effect size.


## Thank you.

## Questions?

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Code available online:
https://sites.google.com/site/icml2014smi/

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[^0]:    ${ }^{1}$ In this talk we use natural logarithms.

